

Bayesian Decision Theory

Berlin Chen 2005

References:

1. E. Alpaydin, *Introduction to Machine Learning*, Chapter 3
2. Tom M. Mitchell, *Machine Learning*, Chapter 6

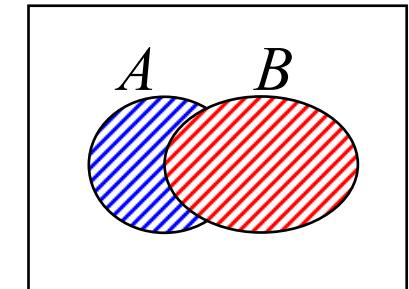
Review: Basic Formulas for Probabilities

- **Product Rule:** probability $P(A \wedge B)$ of a conjunction of two events A and B

$$P(A \wedge B) = P(A, B) = P(A|B)P(B) = P(B|A)P(A)$$

- **Sum Rule:** probability of a disjunction of two events A and B

$$P(A \vee B) = P(A) + P(B) - P(A \wedge B)$$

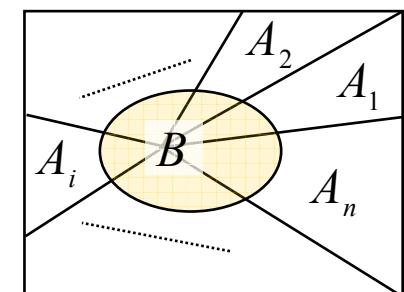


- **Theorem of total probability:** if events A_1, \dots, A_n are **mutually exclusive and exhaustive**

$$\left(\forall i \neq j \quad P(A_i \wedge A_j) = 0 \text{ and } \sum_{i=1}^n P(A_i) = 1 \right)$$

$$P(B) = \sum_{i=1}^n P(B \wedge A_i)$$

$$= \sum_{i=1}^n P(B | A_i) P(A_i)$$



Review: Basic Formulas for Probabilities (cont.)

- **Chain Rule:** probability of a conjunction of many events A_1, A_2, \dots, A_n

$$\begin{aligned} & P(A_1, A_2, \dots, A_n) \\ &= P(A_1)P(A_2 | A_1)P(A_3 | A_1, A_2) \dots P(A_n | A_1, A_2, \dots, A_{n-1}) \end{aligned}$$

Classification

- Illustrative Case 1: Credit Scoring

$$\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}^t = \begin{bmatrix} \text{income} \\ \text{savings} \end{bmatrix}^t$$

$C_1 \Rightarrow$ high - risk

$C_2 = 0 \Rightarrow$ Low - risk

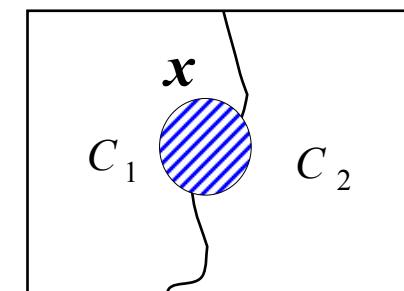
- Given a new application $\mathbf{x} = [x_1, x_2]$

Choose $\begin{cases} C_1 & \text{if } P(C_1|\mathbf{x}) > 0.5 \\ C_2 & \text{otherwise} \end{cases}$

or equivalently

Choose $\begin{cases} C_1 & \text{if } P(C_1|\mathbf{x}) > P(C_2|\mathbf{x}) \\ C_2 & \text{otherwise} \end{cases}$

Note that $P(C_1|\mathbf{x}) + P(C_2|\mathbf{x}) = 1$



Classification (cont.)

- Bayes' Classifier
 - We can use the probability theory to make inference from data

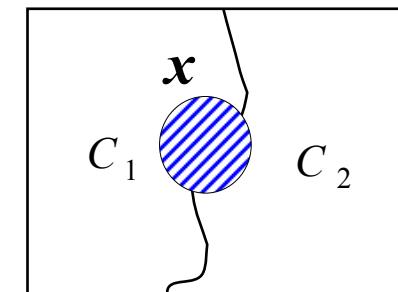
$$P(C|x) = \frac{P(x|C)P(C)}{P(x)}$$

- x : observed data (variable)
- C : class hypothesis
- $P(x)$: prior probability of x
- $P(C)$: prior probability of C
- $P(x|C)$ probability of x given C

Classification (cont.)

- Calculate the posterior probability of the concept C (C_1 or C_2) after having the observation x
 - Combine the prior and what the data tells using Bayes' rule


$$P(C|x) = \frac{P(x|C)P(C)}{P(x)}$$
$$\text{posterior} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$



- C_1 and C_2 are **mutually exclusive and exhaustive** classes (concepts)

$$\begin{aligned}\therefore P(x) &= P(x \wedge C_1) + P(x \wedge C_2) \\ &= P(x|C_1)P(C_1) + P(x|C_2)P(C_2)\end{aligned}$$

Classification (cont.)

- Bayes' Classifier: Extended to K mutually exclusive and exhaustive classes

$$P(C_i) \geq 0 \quad \text{and} \quad \sum_{i=1}^K P(C_i) = 1$$

$$P(C_i | \mathbf{x}) = \frac{P(\mathbf{x}|C_i)P(C_i)}{P(\mathbf{x})} = \frac{P(\mathbf{x}|C_i)P(C_i)}{\sum_{k=1}^K P(\mathbf{x}|C_k)P(C_k)}$$

Classification (cont.)

- Maximum likelihood classifier

- The posterior probability that the data belongs to class C_i

$$L_i(x) = P(x|C_i)$$

- Have the same classification result as that of Bayes' Classifier, if the prior probability $P(C_i)$ is assumed to be equal to each other

$$\max_i L_i(x) = \max_i P(x|C_i) \quad ? \quad \max_i P(C_i|x) = \max_i \frac{P(x|C_i)P(C_i)}{P(x)}$$

Classification: Illustrative Case 2

- Does a patient have cancer (C_1) or not (C_2) ?
 - A patient takes a lab test, and result would be $x = "+"$ or $x = "-"$
 1. If the result comes back positive ($x = "+"$)
 2. And we also knew that the test returns a correct positive result (+) in only 98% of the cases in which the disease is actually present ($P(+|C_1) = 0.98$) and a correct negative result (-) in only 97% of the cases in which the disease is not present ($P(-|C_2) = 0.97$)

Furthermore, 0.008 of the entire population have this cancer ($P(C_1) = 0.008$)

Classification: Illustrative Case 2 (cont.)

- Bayes' Classifier:

$$\begin{aligned} P(C_1|+) &= \frac{P(+|C_1)P(C_1)}{P(+|C_2)P(C_2) + P(+|C_1)P(C_1)} \\ &= \frac{0.98 \times 0.008}{0.03 \times 0.992 + 0.98 \times 0.008} \\ &\approx \frac{0.0078}{0.298 + 0.0078} \\ &\approx 0.21 \end{aligned} \quad \rightarrow \quad \begin{aligned} P(C_1|+) &\approx 0.21 \\ P(C_2|+) &\approx 0.79 \quad \checkmark \end{aligned}$$
$$\begin{aligned} P(C_2|+) &= \frac{P(+|C_2)P(C_2)}{P(+|C_2)P(C_2) + P(+|C_1)P(C_1)} \\ &= \frac{0.03 \times 0.992}{0.03 \times 0.992 + 0.98 \times 0.008} \\ &\approx \frac{0.298}{0.298 + 0.0078} \\ &\approx 0.79 \end{aligned}$$

Classification: Illustrative Case 2 (cont.)

- Maximum likelihood classifier:

$$P(+ | C_1) = 0.98 \quad \checkmark$$

$$P(+ | C_2) = 0.03$$

Losses and Risks

- Decisions are not always perfect
 - E.g., “Loan Application”
 - The loss for a high-risk applicant erroneously accepted (false acceptance) may be different from that for an erroneously rejected low-risk applicant (false rejection)
 - Much critical for other cases such as medical diagnosis or earthquake prediction

Expected Risk

- Def: the **Expected Risk** for taking action α_i
 - Hypothesize that the example x belongs to class i
 - Suppose the example actually belongs to some class k

$$R(\alpha_i | x) = \sum_{k=1}^K \lambda_{ik} P(C_k | x)$$

$$\lambda_{ik} = \begin{cases} 0 & \text{if } i = k \\ 1 & \text{if } i \neq k \end{cases}$$

- A zero-one loss function
- All correct decisions have no loss and all error are equally costly
- Choose the action α_i with minimum risk

$$\alpha = \arg \min_j R(\alpha_j | x)$$

Expected Risk (cont.)

- Choosing the action α_i with minimum risk is equivalent to choosing the class C_i with the highest posterior probability

$$\begin{aligned} R(\alpha_i | \mathbf{x}) &= \sum_{k=1}^K \lambda_{ik} P(C_k | \mathbf{x}) \\ &= \sum_{k=1, k \neq i}^K P(C_k | \mathbf{x}) \\ &= 1 - P(C_i | \mathbf{x}) \end{aligned}$$

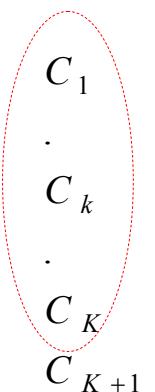
- Choose the action α_i with

$$\alpha_i = \arg \max_j P(C_j | \mathbf{x})$$

Expected Risk: Reject Action Involved

- Manual Decisions ?
 - Wrong decisions (misclassification) may have very high cost
 - Resort to a manual decision when automatic system has low certainty of its decision
- Define an additional action of reject (or doubt) α_{K+1}

$$\lambda_{ik} = \begin{cases} 0 & \text{if } i = k \\ \lambda & \text{if } i = K + 1 \\ 1 & \text{otherwise} \end{cases} \quad \mathbf{x} \rightarrow c_i$$



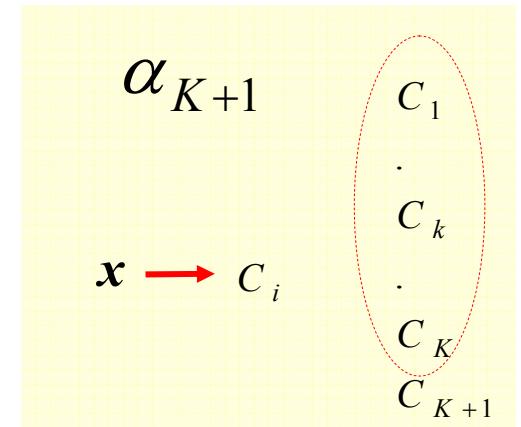
- λ ($0 \leq \lambda \leq 1$) is the loss incurred for choosing the $(K+1)$ st action of reject

Expected Risk: Reject Action Involved (cont.)

- The risk for choosing the reject ((K+1)st) action α_{K+1}

$$R(\alpha_{K+1} | \mathbf{x}) = \sum_{k=1}^K \lambda P(C_k | \mathbf{x})$$

$$\begin{aligned}\alpha_i &= \alpha_{K+1} \\ &= \lambda \sum_{k=1}^K P(C_k | \mathbf{x}) \\ &= \lambda\end{aligned}$$



- Recall that the risk for choosing action α_i ($1 \leq i \leq K$)

$$\begin{aligned}R(\alpha_i | \mathbf{x}) &= \sum_{k=1}^K \lambda_{ik} P(C_k | \mathbf{x}) \\ &= \sum_{k=1, k \neq i}^K P(C_k | \mathbf{x}) \\ &= 1 - P(C_i | \mathbf{x})\end{aligned}$$

Expected Risk: Reject Action Involved (cont.)

- The optimal decision rule is to:

Choose C_i if $R(\alpha_i | \mathbf{x}) < R(\alpha_j | \mathbf{x})$ for all $1 \leq j \leq K, j \neq i$

and $R(\alpha_i | \mathbf{x}) < R(\alpha_{K+1} | \mathbf{x})$

Reject if $R(\alpha_{K+1} | \mathbf{x}) < R(\alpha_j | \mathbf{x})$ for all $1 \leq j \leq K$

- That is

Choose C_i if $P(C_i | \mathbf{x}) > P(C_j | \mathbf{x})$ for all $1 \leq j \leq K, j \neq i$

and $P(C_i | \mathbf{x}) > 1 - \lambda$

Reject otherwise

- When $\lambda = 0 \rightarrow$ always reject the chosen action
- When $\lambda = 1 \rightarrow$ always accept the chosen action

Discriminant Functions

- Classification can be thought of as a set of discriminant functions, $g_i(\mathbf{x}), i = 1, \dots, K$, for each class such that
 - Choose C_i if $g_i(\mathbf{x}) = \max_k g_k(\mathbf{x})$
- $g_i(\mathbf{x})$ can be expressed by using the Bayes's classifier (with minimum risk and no additional action of reject)

$$g_i(\mathbf{x}) = -R(\alpha_i | \mathbf{x})$$

- If the zero-one loss function is imposed, $g_i(\mathbf{x})$ can also be expressed by

$$g_i(\mathbf{x}) = P(C_i | \mathbf{x})$$

- With the same ranking result, we can have $g_i(\mathbf{x}) = P(\mathbf{x} | C_i)P(C_i)$

Discriminant Functions (cont.)

- The instance space thus can be divided into K decision regions R_1, \dots, R_K , where

$$R_i = \left\{ \mathbf{x} \mid g_i(\mathbf{x}) = \max_k g_k(\mathbf{x}) \right\}$$

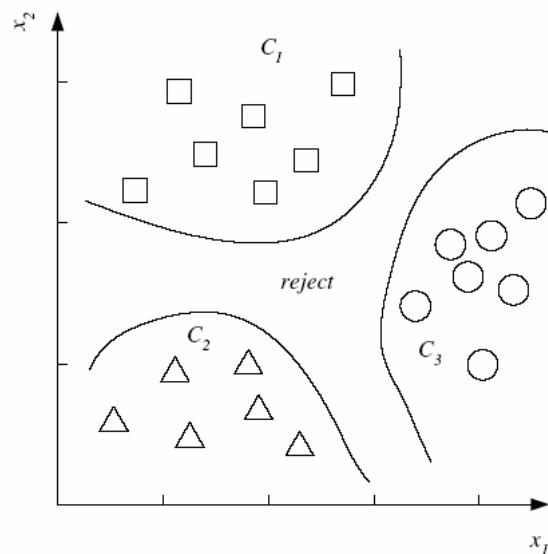


Figure 3.1: Example of decision regions and decision boundaries.

Discriminant Functions (cont.)

- For two-class problems, we can merely define a single discriminant function

$$g(\mathbf{x}) = g_1(\mathbf{x}) - g_2(\mathbf{x})$$

Choose $\begin{cases} C_1 & \text{if } g(\mathbf{x}) > 0 \\ C_2 & \text{otherwise} \end{cases}$

Choosing Hypotheses*: MAP Criterion

- In machine learning, we are interested in finding the best (most probable) hypothesis (classifier) h_{c_i} from some hypothesis space H , given the observed training data set $X_{c_i} = \{(x^t, r^t) | r^t = c_i, t = 1, 2, \dots, n\}$

$$\begin{aligned} h_{MAP} &= \arg \max_{h_{c_i} \in H} P(h_{c_i} | X_{c_i}) \\ &= \arg \max_{h_{c_i} \in H} \frac{P(X_{c_i} | h_{c_i}) P(h_{c_i})}{P(X_{c_i})} \\ &= \arg \max_{h_{c_i} \in H} P(X_{c_i} | h_{c_i}) P(h_{c_i}) \end{aligned}$$

- A Maximum a Posteriori (MAP) hypothesis h_{MAP}

Choosing Hypotheses*: *ML* Criterion

- If we further assume that every hypothesis is equally probable a priori, e.g. $P(h_{c_i}) = P(h_{c_j})$. The above equation can be simplified as:

$$h_{ML} = \arg \max_{h_{c_i} \in H} P(X_{c_i} | h_{c_i})$$

- A Maximum Likelihood (*ML*) hypothesis h_{ML}
- $P(X_{c_i} | h_{c_i})$ often called “the likelihood of the data set X_{c_i} given h_{c_i} ”

Naïve Bayes Classifier

- A simplified approach to the Bayes's classifier
 - The attributes x_1, x_2, \dots, x_d of an instance/example \mathbf{x} are assumed to be independent conditioned on a given class hypothesis
 - Naïve Bayes assumption:

$$P(\mathbf{x}|C_j) = P(x_1, x_2, \dots, x_d | C_j) = \prod_{n=1}^d P(x_n | C_j)$$

- Naïve Bayes Classifier:

$$\begin{aligned} C_{MAP} &= \arg \max_{C_k} P(C_k | \mathbf{x}) \\ &= \arg \max_{C_k} \frac{P(\mathbf{x}|C_k)P(C_k)}{P(\mathbf{x})} = \arg \max_{C_k} P(\mathbf{x}|C_k)P(C_k) \\ &= \arg \max_{C_k} P(x_1, \dots, x_d | C_k)P(C_k) \\ &= \arg \max_{C_k} P(C_k) \prod_{n=1}^d P(x_n | C_k) \end{aligned}$$

Naïve Bayes Classifier (cont.)

$$x = (x_A, x_B, x_C)$$

- Illustrative case 1

- Given a data set with 3-dimensional Boolean examples $x = (x_A, x_B, x_C)$, train a naïve Bayes classifier to predict the classification

Attribute A	Attribute B	Attribute C	Classification D
F	T	F	T
F	F	T	T
T	F	F	T
T	F	F	F
F	T	T	F
F	F	T	F

$$P(D = T) = 1/2, P(D = F) = 1/2$$

$$P(A = T | D = T) = 1/3, P(A = F | D = T) = 2/3$$

$$P(B = T | D = T) = 1/3, P(B = F | D = T) = 2/3$$

$$P(C = T | D = T) = 1/3, P(C = F | D = T) = 2/3$$

$$P(A = T | D = F) = 1/3, P(A = F | D = F) = 2/3$$

$$P(B = T | D = F) = 1/3, P(B = F | D = F) = 2/3$$

$$P(C = T | D = F) = 2/3, P(C = F | D = F) = 1/3$$

- What is the predicted probability $P(D = T | A = T, B = F, C = T)$?
- What is the predicted probability $P(D = T | B = T)$?

Naïve Bayes Classifier (cont.)

- Illustrative case 1 (cont.)

$$(i) P(D = T | A = T, B = F, C = T)$$

$$= \frac{P(A = T, B = F, C = T | D = T)P(D = T)}{P(A = T, B = F, C = T)}$$

$$= \frac{P(A = T, B = F, C = T | D = T)P(D = T)}{P(A = T, B = F, C = T | D = T)P(D = T) + P(A = T, B = F, C = T | D = F)P(D = F)}$$

$$= \frac{\frac{1}{3} \cdot \frac{2}{3} \cdot \frac{1}{3} \cdot \frac{1}{2}}{\frac{1}{3} \cdot \frac{2}{3} \cdot \frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{2}{3} \cdot \frac{2}{3} \cdot \frac{1}{2}} = \frac{2}{2+4} = \frac{1}{3}$$

$$(ii) P(D = T | B = T)$$

$$= \frac{P(B = T | D = T)P(D = T)}{P(B = T)}$$

$$= \frac{P(B = T | D = T)P(D = T)}{P(B = T | D = T)P(D = T) + P(B = T | D = F)P(D = F)}$$

$$= \frac{\frac{1}{3} \cdot \frac{1}{2}}{\frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2}} = \frac{1}{1+1} = \frac{1}{2}$$

How to Train a Naïve Bayes Classifier

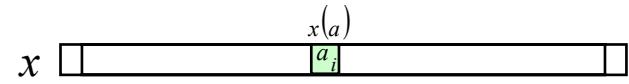
- `Naïve_Bayes_Learn(examples)`

For each target value v_j

$$\hat{P}(v_j) \leftarrow \text{maximum likelihood (ML) estimate of } P(v_j)$$

For each attribute value a_i of each attribute a

$$\hat{P}(a_i | v_j) \leftarrow \text{maximum likelihood (ML) estimate of } P(a_i | v_j)$$



$$\frac{|v_j|}{n} \text{ or } \frac{|v_j|}{\sum_{v_k} |v_k|}$$

$$\frac{\sum_{x \in v_j, x(a)=a_i} 1}{\sum_{x \in v_j} 1} = \frac{\sum_{x \in v_j, x(a)=a_i} 1}{|v_j|}$$

- `Classify_New_Instance(x)`

$$v_{NB} = \arg \max_{v_j \in V} P(v_j) \prod_{a_i \in x} P(a_i | v_j)$$

Naïve Bayes: Example 2

- Consider *PlayTennis* again and new instance
 $\langle Outlook=sunny, Temperature=cool, Humidity=high, Wind=strong \rangle$
- Want to compute

$$v_{NB} = \arg \max_{v_j \in V = \{yes, no\}} P(v_j) \times P(Outlook=sunny|v_j) \times P(Temperature=cool|v_j) \\ \times P(Humidity=high|v_j) \times P(Wind=Strong|v_j)$$

$$P(yes) \times P(Outlook=sunny|yes) \times P(Temperature=cool|yes) \\ \times P(Humidity=high|yes) \times P(Wind=Strong|yes) = 0.0053$$

$$P(no) \times P(Outlook=sunny|no) \times P(Temperature=cool|no) \\ \times P(Humidity=high|no) \times P(Wind=Strong|no) = 0.206$$

$$\therefore v_{NB} = no$$

Dealing with Data Sparseness

- What if none of the training instances with target value v_j have attribute value a_i ? Then

$$\hat{P}(a_i | v_j) = 0, \text{ and } \dots$$

$$v_{NB} = \arg \max_{v_j \in V} \hat{P}(v_j) \prod_i \hat{P}(a_i | v_j)$$

- Typical solution is Bayesian estimate for $\hat{P}(a_i | v_j)$

$$\hat{P}(a_i | v_j) \leftarrow \frac{n_c + mp}{n + m} \quad \text{Smoothing}$$

- n is number of training examples for which $v = v_j$
- n_c is number of training examples for which $v = v_j$ and $a = a_i$
- p is prior estimate for $\hat{P}(a_i | v_j)$
- m is weight given to prior (i.e., number of “virtual” examples)

Example: Learning to Classify Text

- For instance,
 - Learn which news articles are of interest
 - Learn to classify web pages by topic
- Naïve Bayes is among the most effective algorithms
- What attributes shall we use to represent text documents
 - The word occurs in each document position

Example: Learning to Classify Text (cont.)

- Target Concept: *Interesting* ? Document $\rightarrow \{+, -\}$
 1. Represent each document by vector of words
 - one attribute per word position in document
 2. Learning Use training examples to estimate
 - $P(+)$
 - $P(-)$
 - $P(doc|+)$
 - $P(doc|-)$
- Naïve Bayes conditional independence assumption

$$P(doc|v_j) = \prod_{i=1}^{\text{length}(doc)} P(a_i = w_k | v_j)$$

- Where $P(a_i = w_k | v_j)$ is probability that word in position i is w_k , given v_j
- One more assumption: $P(a_i = w_k | v_j) = P(a_m = w_k | v_j) \forall i, m$ Time Invariant

$doc(i) = a_i = w_k$
 doc 

Example: Learning to Classify Text (cont.)

- Learn_Naïve_Bayes_Text(Examples, V)
 1. Collect all words and other tokens that occur in *Examples*
 - $Vocabulary \leftarrow$ all distinct words and other tokens in Examples
 2. Calculate the required $P(v_j)$ and $P(w_k | v_j)$ probability terms
 - $docs_j \leftarrow$ subset of Examples for which the target value is v_j
 - $P(v_j) \leftarrow \frac{|docs_j|}{|Examples|}$
 - $Text_j \leftarrow$ a single document created by concatenating all members of $docs_j$
 - $n \leftarrow$ total number of words in $Text_j$ (counting duplicate words multiple times)
 - For each word w_k in $Vocabulary$
 - $n_k \leftarrow$ number of times word w_k occurs in $Text_j$
 - $P(w_k | v_j) \leftarrow \frac{n_k + 1}{n + |Vocabulary|}$ Smoothed unigram

Example: Learning to Classify Text (cont.)

- **Classify_Naïve_Bayes_Text(*Doc*)**
 - *positions* \leftarrow all word positions in *Doc* that contain tokens found in *Vocabulary*
 - Return v_{NB} , where

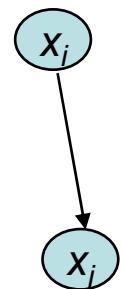
$$v_{NB} = \arg \max_{v_j \in V} P(v_j) \prod_{i \in positions} P(a_i | v_j)$$

Bayesian Networks

- Premise
 - Naïve Bayes assumption of conditional independence too restrictive
 - But it is intractable without some such assumptions
 - Bayesian networks describe conditional independence among subsets of variables
 - Allows combining prior knowledge about (in)dependencies among variables with observed training data
- Bayesian Networks also called
 - Bayesian Belief Networks, Bayes Nets, Belief Networks, Probabilistic Networks, Graphical Models etc.

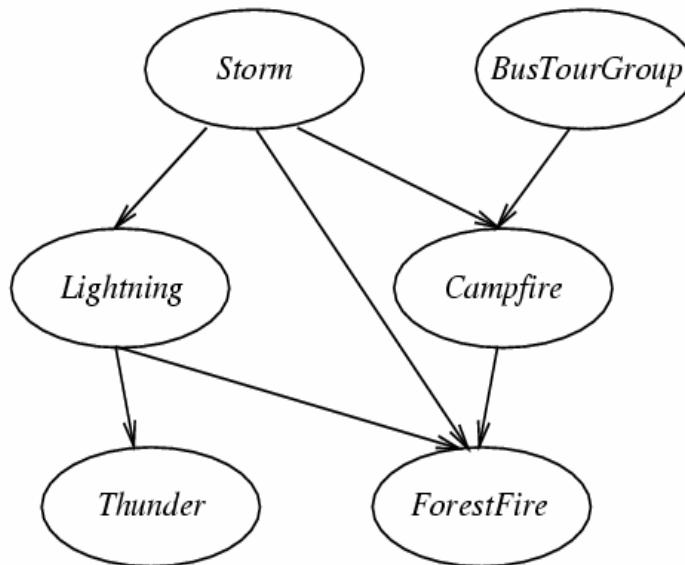
Bayesian Networks (cont.)

- A simple, graphical notation for conditional independence assertions and hence for compact specification of full joint distributions
- **Syntax**
 - A set of nodes, one per variable (**discrete or continuous**)
 - **For discrete variable, they can be either binary or not**
 - A directed, acyclic graph (link/arrow \approx “directly influences”)
 - A conditional distribution for each node given its parents
$$P(X_i | Parents(X_i))$$
- In the simplest case, conditional distribution represented as a Conditional Probability Table (CPT) giving the distribution over $P(X_i)$ for each combination of parent values



Bayesian Networks (cont.)

- E.g., nodes of discrete binary variables



Conditional Probability Table (CPT)		
S	B	P(C)
T	T	0.4
T	F	0.1
F	T	0.8
F	F	0.2

Campfire

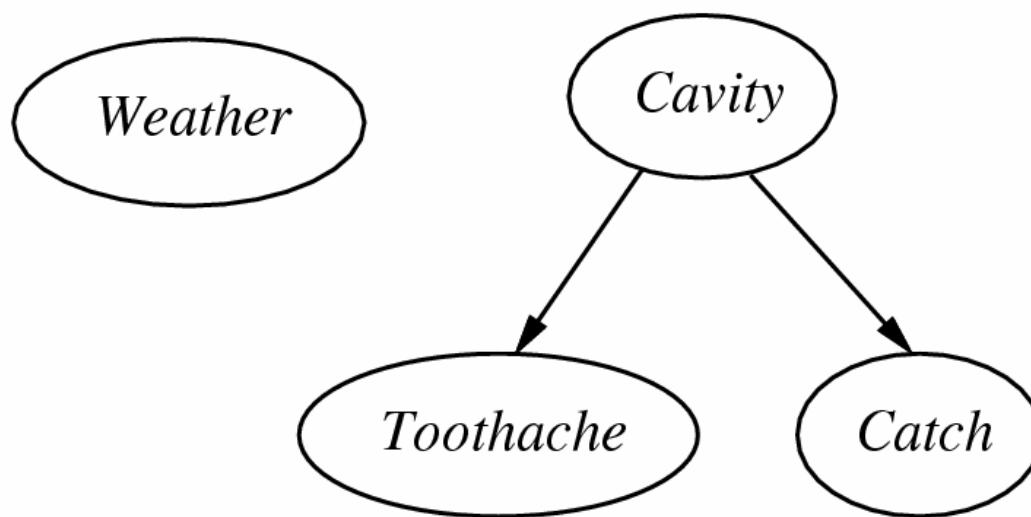
A conditional probability table (CPT) for the node *Campfire* is shown. The table has four columns corresponding to the combinations of parents *S* and *B*: S, B , $S, \neg B$, $\neg S, B$, and $\neg S, \neg B$. The rows are labeled by the values of *Campfire*: C and $\neg C$. The entries in the table are as follows:

S, B	$S, \neg B$	$\neg S, B$	$\neg S, \neg B$	
C	0.4	0.1	0.8	0.2
$\neg C$	0.6	0.9	0.2	0.8

- Each node is asserted to be conditionally independent of its nondescendants, given its immediate predecessors
- Directed acyclic graph

Example 1:Dentist Network

- Topology of network encodes conditional independence assertions



- *Weather* is independent of the other variables
- *Toothache* and *Catch* are conditionally independent given *Cavity*
 - *Cavity* is a direct cause of *Toothache* and *Catch*

Conditional (In)dependence

- Definition: X is conditionally independent of Y given Z if the probability distribution governing X is independent of the value of Y given the value of Z ; that is, if

$$(\forall x_i, y_j, z_k) P(X = x_i | Y = y_j, Z = z_k) = P(X = x_i | Z = z_k)$$

- More compactly, we can write

$$P(X|Y, Z) = P(X|Z)$$

- Conditional independence allows breaking down inference into calculation over small group of variables

Conditional (In)dependence (cont.)

- Example: *Thunder* is conditionally independent of *Rain* given *Lightning*

$$P(\text{Thunder}|\text{Rain}, \text{Lightning}) = P(\text{Thunder}|\text{Lightning})$$

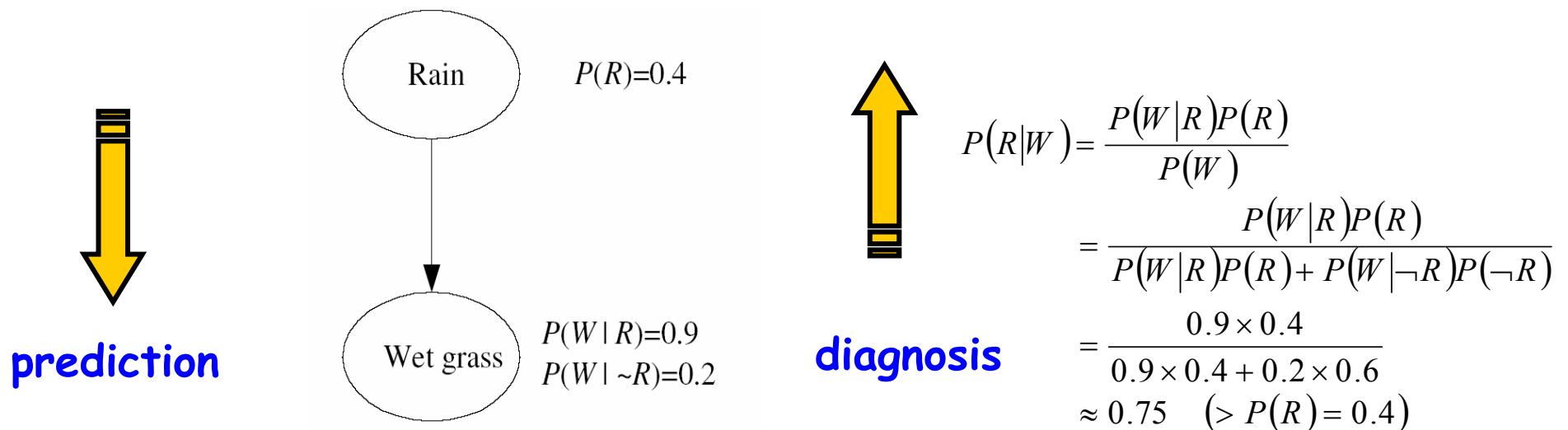
- Recall that Naïve Bayes uses conditional independence to justify

$$P(X, Y|Z) = P(X|Y, Z)P(Y|Z) = P(X|Z)P(Y|Z)$$

X, Y are mutually
independent given Z

Conditional (In)dependence (cont.)

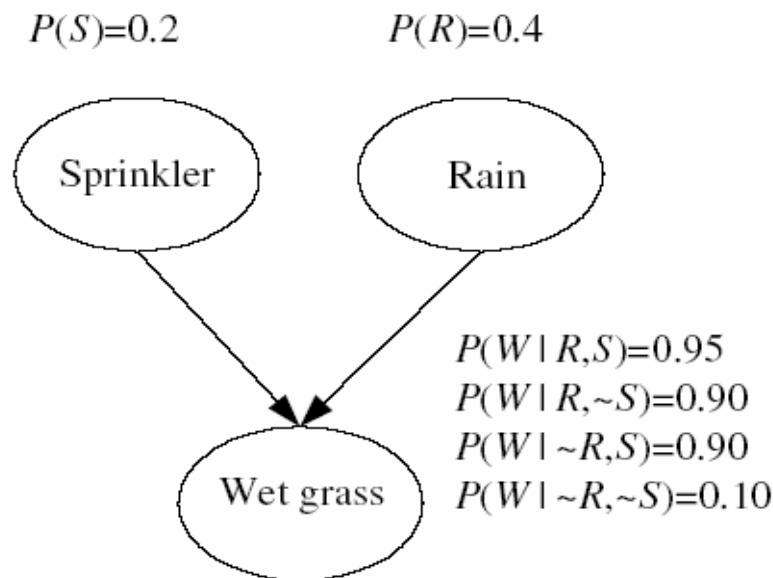
- Bayesian Network also can be thought of as a causal graph in that illustrates causalities between variables
 - We can make a diagnostic inference from the it $P(R|W) = ?$



Conditional (In)dependence (cont.)

- Suppose that sprinkler is included as another cause of wet grass

- Predictive inference



$$\begin{aligned}
 P(W) &= P(W|R, S)P(R, S) + P(W|\neg R, S)P(\neg R, S) \\
 &\quad + P(W|R, \neg S)P(R, \neg S) + P(W|\neg R, \neg S)P(\neg R, \neg S) \\
 &= P(W|R, S)P(R)P(S) + P(W|\neg R, S)P(\neg R)P(S) \\
 &\quad + P(W|R, \neg S)P(R)P(\neg S) + P(W|\neg R, \neg S)P(\neg R)P(\neg S) \\
 &= 0.95 \times 0.4 \times 0.2 + 0.9 \times 0.6 \times 0.2 + 0.9 \times 0.4 \times 0.8 + 0.1 \times 0.6 \times 0.8 \\
 &= 0.52
 \end{aligned}$$

$$\begin{aligned}
 P(W|S) &= P(W|R, S)P(R|S) + P(W|\neg R, S)P(\neg R|S) \\
 &= P(W|R, S)P(R) + P(W|\neg R, S)P(\neg R) \\
 &= 0.95 \times 0.4 + 0.9 \times 0.6 \\
 &\approx 0.92
 \end{aligned}$$

- Diagnostic inference (I)

$$\begin{aligned}
 P(S|W) &= \frac{P(W|S)P(S)}{P(W)} \\
 &\approx \frac{0.92 \times 0.2}{0.52} \\
 &\approx 0.35 \quad (> P(S) = 0.2)
 \end{aligned}$$

Conditional (In)dependence (cont.)

– Diagnostic inference (II)

$$\begin{aligned} & P(S|R, W) \\ &= \frac{P(W|R, S)P(S|R)}{P(W|R)} \\ &= \frac{P(W|R, S)P(S)}{P(W|R)} \\ &= \frac{0.95 \times 0.2}{0.91} \\ &\approx 0.21 \quad (> P(S) = 0.2) \end{aligned}$$

$$\begin{aligned} & P(W|R) \\ &= P(W|R, S)P(S|R) + P(W|R, \neg S)P(\neg S|R) \\ &= P(W|R, S)P(S) + P(W|R, \neg S)P(\neg S) \\ &= 0.95 \times 0.2 + 0.9 \times 0.8 \\ &= 0.91 \end{aligned}$$

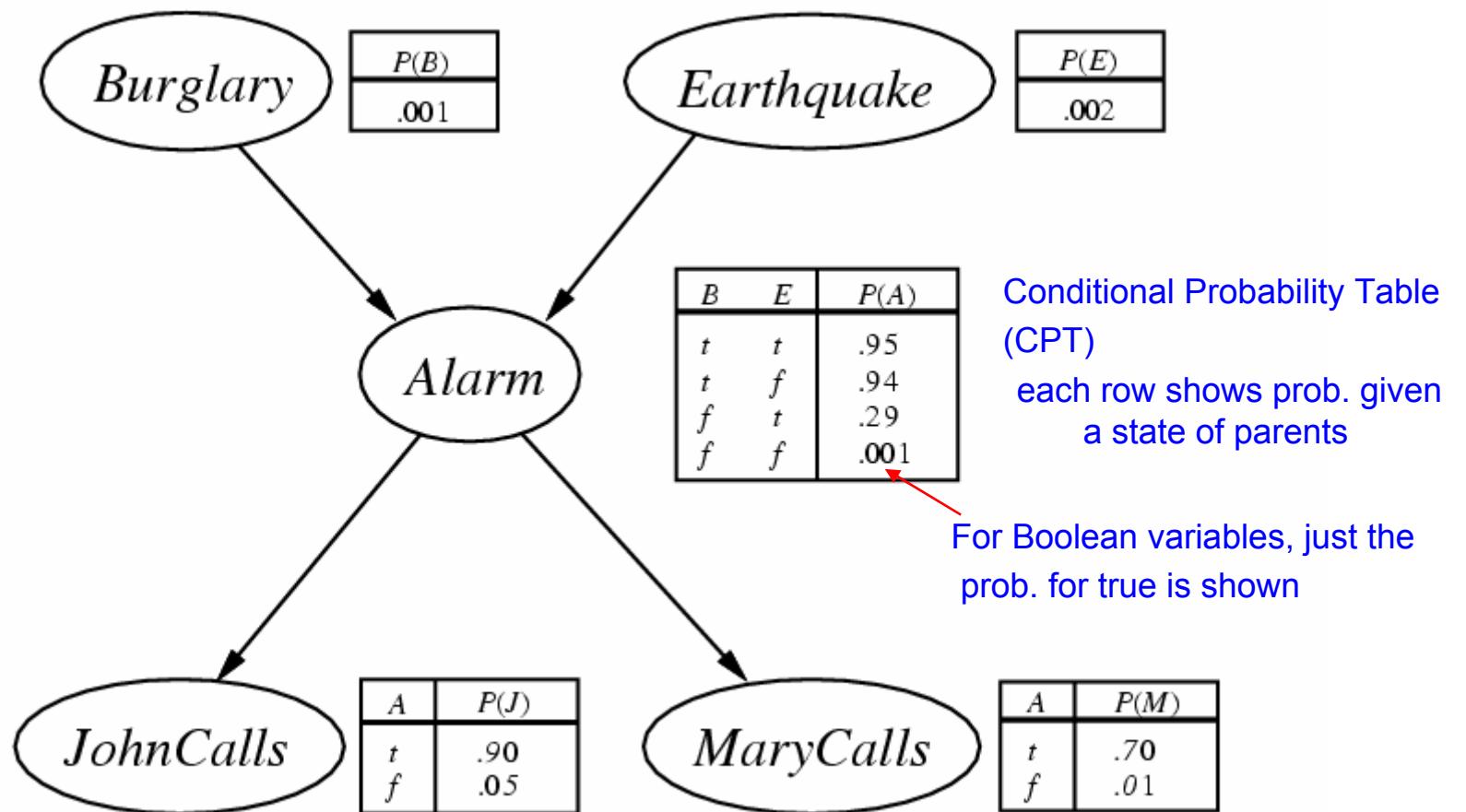
Example 2: Burglary Network

- You're at work, neighbor John calls to say your alarm is ringing, but neighbor Mary doesn't call. Sometimes it's set off by minor earthquakes. Is there a burglar?

$$P(\text{Burglary} = T | \text{JohnCall} = T, \text{MaryCall} = F)$$

- Variables: `Burglar`, `Earthquake`, `Alarm`, `JohnCalls`, `MaryCalls`
- Network topology reflects “causal” knowledge
 - A burglar can set the alarm off
 - An earthquake can set the alarm off
 - The alarm can cause Mary to call
 - The alarm can cause John to call
- But
 - John sometimes confuses the telephone ringing with the alarm
 - Mary likes rather loud music and sometimes misses the alarm

Example 2: Burglary Network



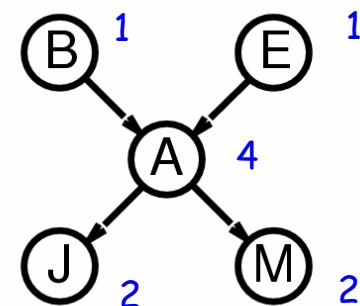
Compactness

- A CPT for Boolean X_i with k Boolean (true/false) parents has 2^k rows for the combinations of parent values
- Each row requires one number p for $X_i = \text{true}$ (the number for $X_i = \text{false}$ is just $1-p$)
- If each variable has no more than k parents, the complete network requires $O(n \cdot 2^k)$ numbers
 - I.e., grows linearly with n , vs. $O(2^n)$ for the full joint distribution
- For burglary net, $1 + 1 + 4 + 2 + 2 = 10$ numbers
 (vs. $2^5 - 1 = 31?$)

$$2^0 \quad 2^0 \quad 2^2 \quad 2^1 \quad 2^1$$

Chain rule

$$\begin{aligned}
 & P(B, E, A, J, M) \\
 &= P(B)P(E|B)P(A|B, E)P(J|B, E, A)P(M|B, E, A, J) \\
 &\approx P(B)P(E)P(A|B, E)P(J|A)P(M|A)
 \end{aligned}$$



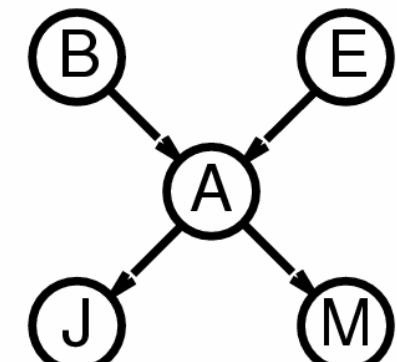
Global Semantics

- Global semantics defines the **full joint distribution** as the product of the local conditional distributions

$$P(X_1, \dots, X_n) \approx \prod_{i=1}^n P(X_i | \text{Parents}(X_i))$$

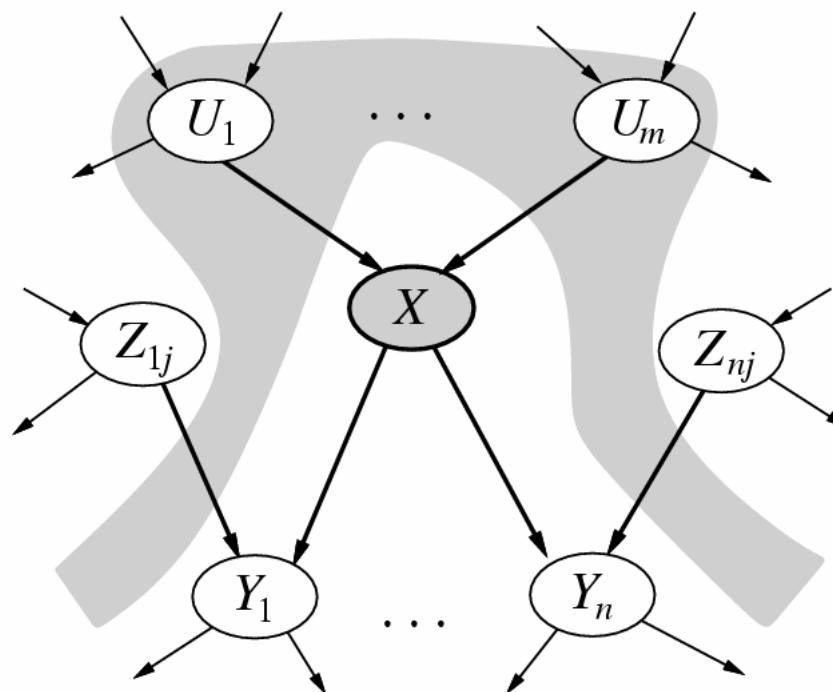
- The Bayesian Network is **semantically**
 - A representation of the joint distribution
 - A encoding of a collection of conditional independence statements
- E.g.,

$$\begin{aligned} & P(J \wedge M \wedge A \wedge \neg B \wedge \neg E) \\ & \approx P(J|A)P(M|A)P(A|\neg B \wedge \neg E)P(\neg B)P(\neg E) \\ & = 0.90 \times 0.70 \times 0.001 \times 0.999 \times 0.998 \\ & = 0.00062 \end{aligned}$$



Local Semantics

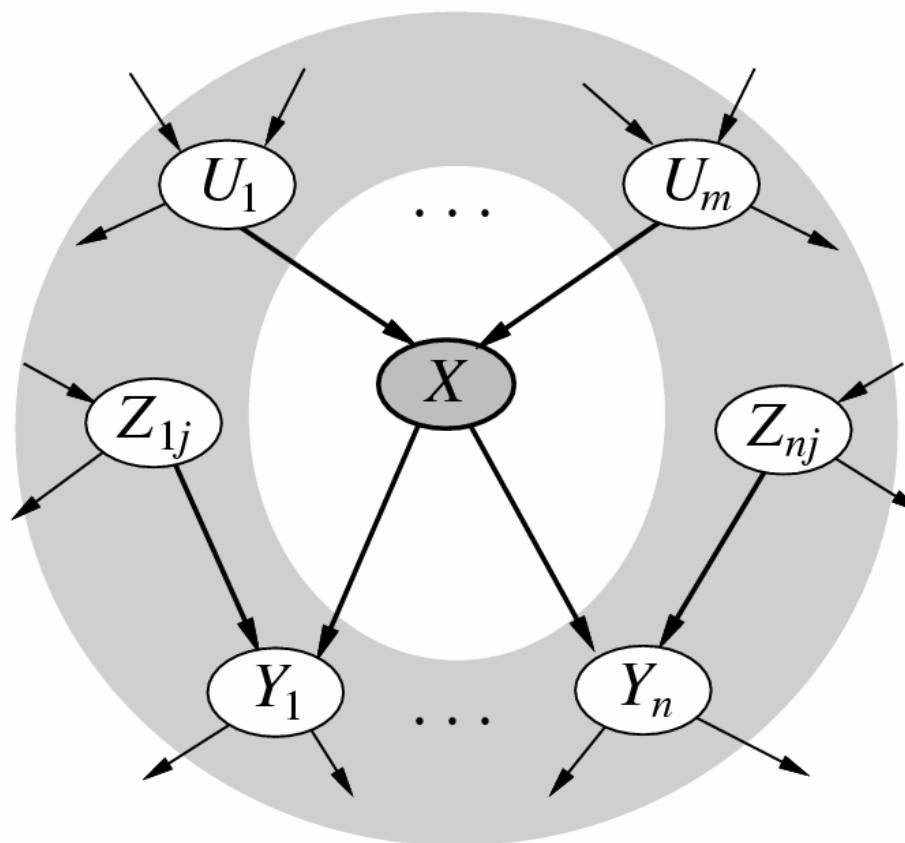
- Local semantics: each node is conditionally independent of its nondescendants **given its parents**



- Local semantics \iff global semantics

Markov Blanket

- Each node is conditionally independent of all others given its parents + children + children's parents



Constructing Bayesian Networks

- Need a method such that a series of locally testable assertions of conditional independence guarantees the required global semantics
 1. Choose an ordering of variables X_1, \dots, X_n
 2. For $i=1$ to n
add X_i to the network and select parents from X_1, \dots, X_{i-1} such that
$$Parents(X_i) \subseteq \{X_1, \dots, X_{i-1}\}$$

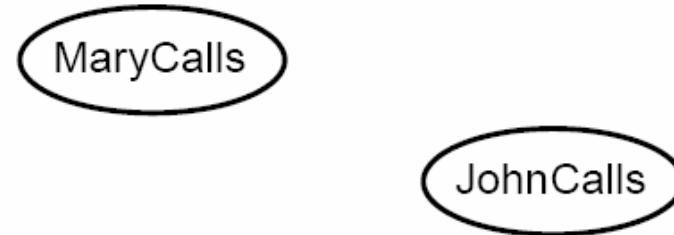
$$P(X_i | X_1, \dots, X_{i-1}) = P(X_i | Parents(X_i))$$

This choice of parents guarantees the global semantics

$$\begin{aligned} P(X_1, \dots, X_n) &= \prod_{i=1}^n P(X_i | X_1, \dots, X_{i-1}) && \text{(chain rule)} \\ &= \prod_{i=1}^n P(X_i | Parents(X_i)) && \text{(by construction)} \end{aligned}$$

Example for Constructing Bayesian Network

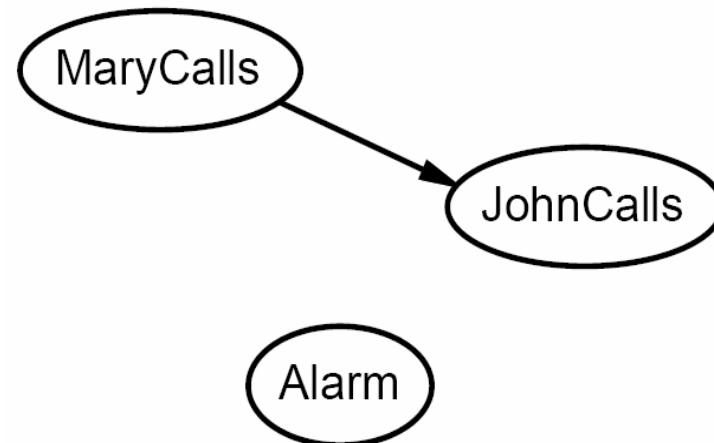
- Suppose we choose the ordering: M, J, A, B, E



- $P(J|M) = P(J)$?

Example (cont.)

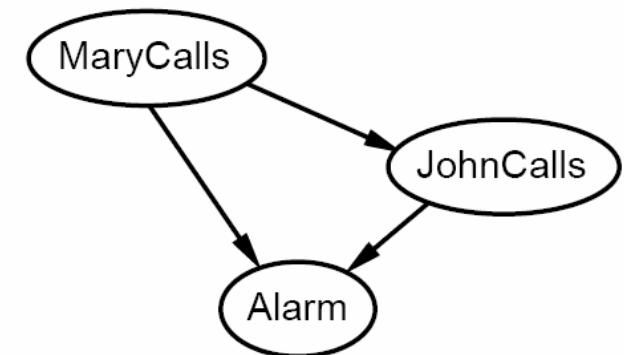
- Suppose we choose the ordering: M, J, A, B, E



- $P(J|M) = P(J)$? **No**
- $P(A|J,M) = P(A|J)$? $P(A|J,M) = P(A)$?

Example (cont.)

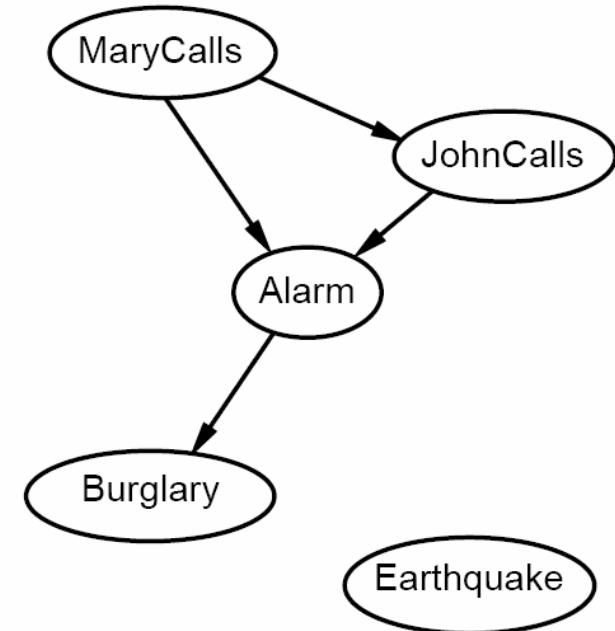
- Suppose we choose the ordering: M, J, A, B, E



- $P(J|M) = P(J)$? **No**
- $P(A|J,M) = P(A|J)$? **No** $P(A|J,M) = P(A)$? **No**
- $P(B|A,J,M) = P(B|A)$?
- $P(B|A,J,M) = P(B)$?

Example (cont.)

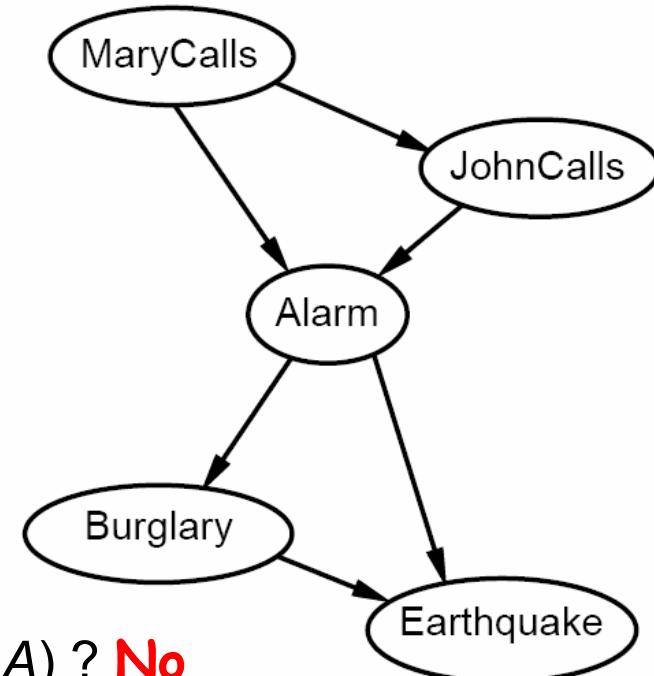
- Suppose we choose the ordering: M, J, A, B, E



- $P(J|M) = P(J)$? **No**
- $P(A|J,M) = P(A|J)$? **No** $P(A|J,M) = P(A)$? **No**
- $P(B|A,J,M) = P(B|A)$? **Yes**
- $P(B|A,J,M) = P(B)$? **No**
- $P(E|B,A,J,M) = P(E|A)$?
- $P(E|B,A,J,M) = P(E|B,A)$?

Example (cont.)

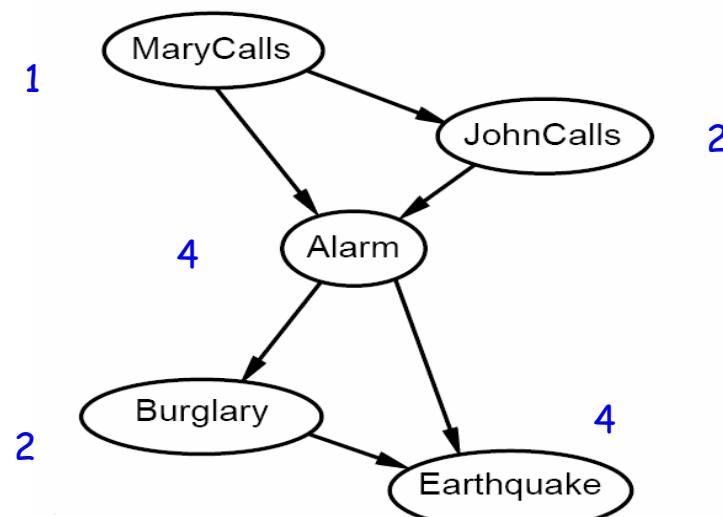
- Suppose we choose the ordering: M, J, A, B, E



- $P(J|M) = P(J)$? **No**
- $P(A|J,M) = P(A|J)$? **No** $P(A|J,M) = P(A)$? **No**
- $P(B|A,J,M) = P(B|A)$? **Yes**
- $P(B|A,J,M) = P(B)$? **No**
- $P(E|B,A,J,M) = P(E|A)$? **No**
- $P(E|B,A,J,M) = P(E|B,A)$? **Yes**

Example (cont.)

- Summary
 - Deciding conditional independence is hard in noncausal directions
(Causal models and conditional independence seem hardwired for humans!)
 - Assessing conditional probabilities is hard in noncausal directions
 - Network is less compact: $1 + 2 + 4 + 2 + 4 = 13$ numbers needed



Inference Tasks

- Simple queries: compute posterior marginal $P(X_i | E = e)$

– E.g.,

$$P(\text{Burglary} | \text{JohnCalls} = \text{true}, \text{MarryCalls} = \text{true})$$

- Conjunctive queries:

$$P(X_i, X_j | E = e) = P(X_i | X_j, E = e)P(X_j | E = e)$$

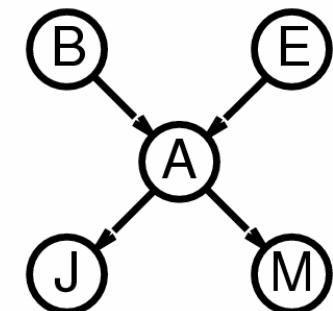
- Optimal decisions: probabilistic inference

$$P(\text{Outcome} | \text{Action}, \text{Evidence})$$

Inference by Enumeration

- Slightly intelligent way to sum out variables from the joint without actually constructing its explicit representation
- Simple query on the burglary network

$$\begin{aligned} P(B|j,m) &= \frac{P(B,j,m)}{P(j,m)} \\ &= \alpha P(B,j,m) \quad (\alpha = 1/P(j,m)) \\ &= \alpha \sum_e \sum_a P(B,e,a,j,m) \end{aligned}$$

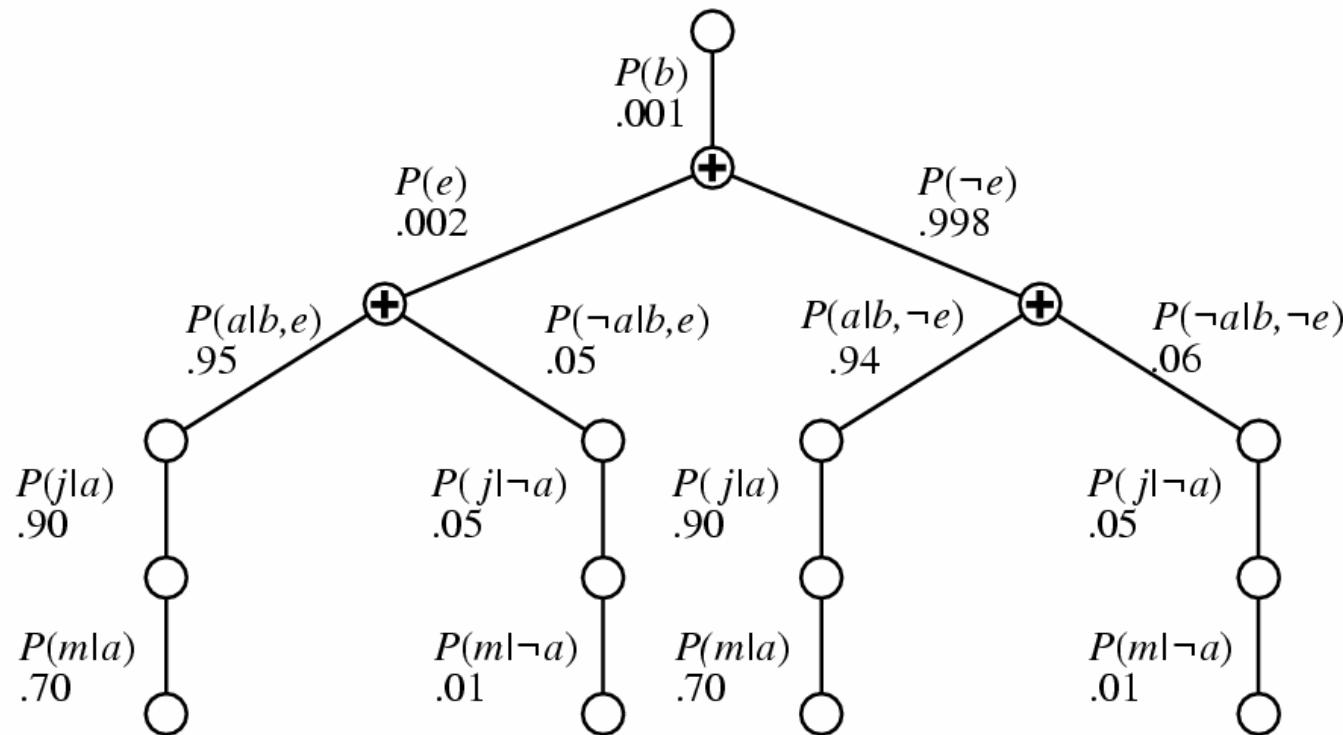


- Rewrite full joint entries using product of CPT entries:

$$\begin{aligned} P(B|j,m) &= \alpha \sum_e \sum_a P(B,e,a,j,m) \\ &= \alpha \sum_e \sum_a P(B)P(e)P(a|B,e)P(j|a)P(m|a) \\ &= \alpha P(B) \sum_e P(e) \sum_a P(a|B,e)P(j|a)P(m|a) \end{aligned}$$

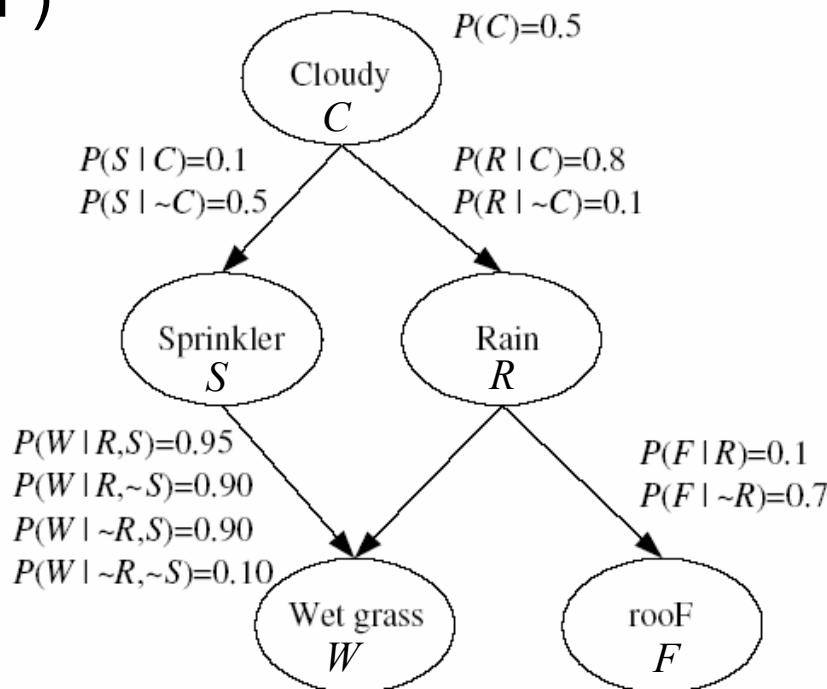
Evaluation Tree

- Enumeration is inefficient: repeated computation\al
 - E.g., computes $P(j|a)P(m|a)$ for each value of e



HW-4: Bayesian Networks

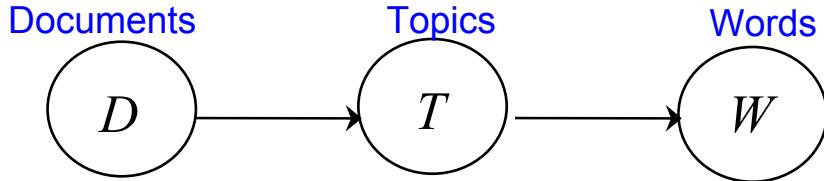
- A new binary variable concerning cat making noise on the roof (**rooF**)



– Predictive inferences

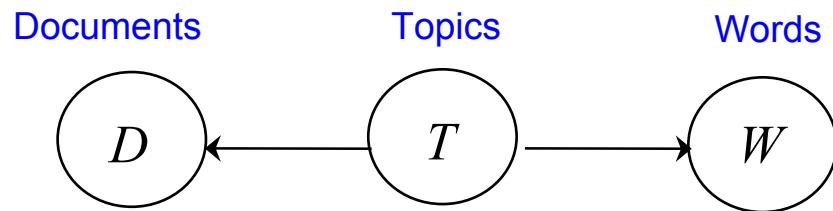
- $P(F | C) = ?$
- $P(F | S) = ?$

Bayesian Networks for Information Retrieval



$$\begin{aligned} P(d_j, w_i) &= \sum_{k=1}^T P(d_j, t_k, w_i) \\ &= \sum_{k=1}^T P(d_j)P(t_k | d_j)P(w_i | t_k) \\ &= P(d_j) \sum_{k=1}^T P(t_k | d_j)P(w_i | t_k) \\ \Rightarrow P(w_i | d_j) &= \frac{P(d_j, w_i)}{P(d_j)} = \sum_{k=1}^T P(t_k | d_j)P(w_i | t_k) \end{aligned}$$

Bayesian Networks for Information Retrieval



$$\begin{aligned} P(d_j, w_i) &= \sum_{k=1}^T P(d_j, t_k, w_i) \\ &= \sum_{k=1}^T P(t_k) P(d_j | t_k) P(w_i | t_k) \\ \Rightarrow P(w_i | d_j) &= \frac{P(d_j, w_i)}{P(d_j)} = \sum_{k=1}^T \frac{P(d_j, t_k)}{P(d_j)} P(w_i | t_k) = \sum_{k=1}^T P(t_k | d_j) P(w_i | t_k) \end{aligned}$$